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Last time:

$W_N(n_1)$  probability of taking  $n_1$  steps to the right after  $N$  steps.

$$\langle n_1 \rangle = Np \quad \langle n_2 \rangle = Nq$$

$$\overline{\Delta n_1^2} = \langle n_1^2 \rangle - \langle n_1 \rangle^2 = Npq \quad \text{and} \quad \Delta^* n_1 = \sqrt{Npq}$$

Now let's find  $\langle m \rangle$  and  $\overline{\Delta m^2}$  for  $P_N(m)$

$$m = n_1 - n_2 \quad N = n_1 + n_2 \therefore m = 2n_1 - N$$

$$\begin{aligned} \langle m \rangle &= 2 \langle n_1 \rangle - N = 2Np - N = N(2p - 1) = \\ &= N(p + \frac{p-1}{-q}) = N(p - q) \end{aligned}$$

probability of being at a distance  $m$  from the start after  $N$  steps

$$\begin{aligned}\Delta m &= m - \langle m \rangle = 2n_1 - N - N(p-q) = 2n_1 - N - N(2p-1) \\ &= 2(n_1 - Np) - N + N = 2(n_1 - \langle n_1 \rangle) = 2\Delta n_1\end{aligned}$$

$$\overline{\Delta m^2} = 4 \overline{\Delta n_1^2} = 4Npq$$

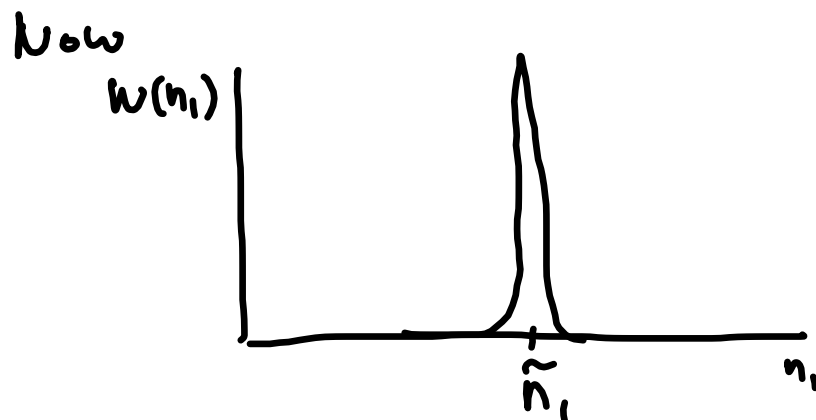
$$\Delta^* m = 2\sqrt{Npq}$$

If  $p=q$  then  $\langle m \rangle = 0$  and  $\Delta^* m = \sqrt{N}$

## Large $N$ limit of the probability distribution.

If  $N$  is very large we'll see that  $W_N(n_i)$  and  $P_N(m)$  become gaussian except for the case in which  $p \ll 1$  or  $q \ll 1$ . Then you'll get a Poisson distribution that does not have a finite dispersion (see Hw #1).

$$\text{If } N \rightarrow \infty \quad \frac{\Delta n_i^*}{\langle n_i \rangle} = \sqrt{\frac{g}{p}} \frac{1}{\sqrt{N}}$$



Since  $w(n_i)$  has a sharp maximum at  $n_i = \tilde{n}_i$ , we can make a Taylor expansion of  $w(n_i)$  about  $\tilde{n}_i$ .

Notice that close to  $\tilde{n}_i$ ,  $|w(n_{i+1}) - w(n_i)| \ll w(n_i)$

so  $n_i$  can be considered continuous.

We know that

$$\left. \frac{\partial W}{\partial n_i} \right|_{\tilde{n}_i} = 0$$

or

$$\left. \frac{d \ln W}{d n_i} \right|_{\tilde{n}_i} = 0$$

A series for  $\ln W$  converges faster than for  $w$ .

$$\ln W(n_1) = \underbrace{\ln W(\tilde{n}_1)}_{\tilde{W}} + \underbrace{\frac{d \ln W}{dn_1}}_0 \Big|_{\tilde{n}_1} \eta + \frac{1}{2} \underbrace{\frac{d^2 \ln W}{dn_1^2}}_{-|B_2|} \Big|_{\tilde{n}_1} \eta^2 + \dots$$

Then

$$W(n_1) \approx \tilde{W} e^{-\frac{1}{2} |B_2| \eta^2} \quad \textcircled{1}$$

Also we know that

$$W_N(n_1) = \frac{N!}{n_1! (N-n_1)!} p^{n_1} q^{N-n_1}$$

$$\ln W_N(n_1) = \ln N! - \ln n_1! - \ln (N-n_1)! + n_1 \ln p + (N-n_1) \ln q \quad \textcircled{2}$$

For large  $n$ ,  $\ln n!$  can be considered continuous.

$$\frac{d \ln n!}{dn} \approx \frac{\ln (n+1)! - \ln n!}{1} = \ln \left[ \frac{(n+1)!}{n!} \right] = \ln(n+1) \approx \ln n \quad \text{since } n \gg 1. \quad (3)$$

Then

$$\frac{d \ln W}{dn_1} \approx -\ln n_1 + \ln(N - n_1) + \ln p - \ln q$$

We can find  $\tilde{n}_1$  since  $\frac{d \ln W}{dn_1} \Big|_{\tilde{n}_1} = 0$

$$0 = \ln \frac{(N - \tilde{n}_1) p}{\tilde{n}_1 q}$$

Then  $N - \tilde{n}_1 = \tilde{n}_1 \frac{q}{p}$

$$N = \tilde{n}_1 \left( \frac{q}{p} + 1 \right) = \tilde{n}_1 \frac{q+p}{p}$$

$\tilde{n}_1 = NP$

Also

$$B_2 = \left. \frac{d^2 \ln W}{du_1^2} \right|_{\tilde{n}_1} = - \frac{1}{n_1} \Big|_{\tilde{n}_1} - \frac{1}{n-n_1} \Big|_{\tilde{n}_1} = - \frac{1}{Np} - \frac{1}{N-Np}$$

$$= - \frac{1}{Np} - \frac{1}{N(1-p)} = - \frac{q+p}{Npq} = - \frac{1}{Npq}$$

$$|B_2| = \frac{1}{Npq} \equiv \frac{1}{\Delta n^2}$$

To find  $\tilde{W}$  we request normalization of  $W(n_i)$ :

$$1 = \int_{-\infty}^{\infty} W(n_i) du_1 \approx \int_{-\infty}^{\infty} W(\tilde{n}_1 + \eta) d\eta = \int_{-\infty}^{\infty} \tilde{W} e^{-\frac{|B_2|}{2} \eta^2} d\eta = 1$$

because  $W(n_i) \approx 0$  for  $n_i$  away from  $\tilde{n}_1$ .

We know that

$$\int_{-\infty}^{\infty} e^{-ax^2} dx = \sqrt{\frac{\pi}{a}}$$

then

$$I = \tilde{W} \sqrt{\frac{2\pi}{|B_2|}} = 1 \quad \Rightarrow \quad \tilde{W} = \sqrt{\frac{|B_2|}{2\pi}} = \sqrt{\frac{1}{2\pi Npq}}$$

Then

$$W(n_1) = \sqrt{\frac{1}{2\pi Npq}} e^{-\frac{1}{2} \frac{(n_1 - Np)^2}{2Npq}}$$

or to remove  $N$ :

$$W(n_1) = \sqrt{\frac{1}{2\pi (\overline{Npq})}} e^{-\frac{1}{2} \frac{(n_1 - \langle n_1 \rangle)^2}{2(\overline{Npq})}}$$

(4)

Gaussian shape.



## Gaussian probability distribution

From (9) we can also obtain the large  $N$  limit of  $P_N(m)$ :

$$m = 2n_1 - N \Rightarrow n_1 = \frac{m+N}{2}$$

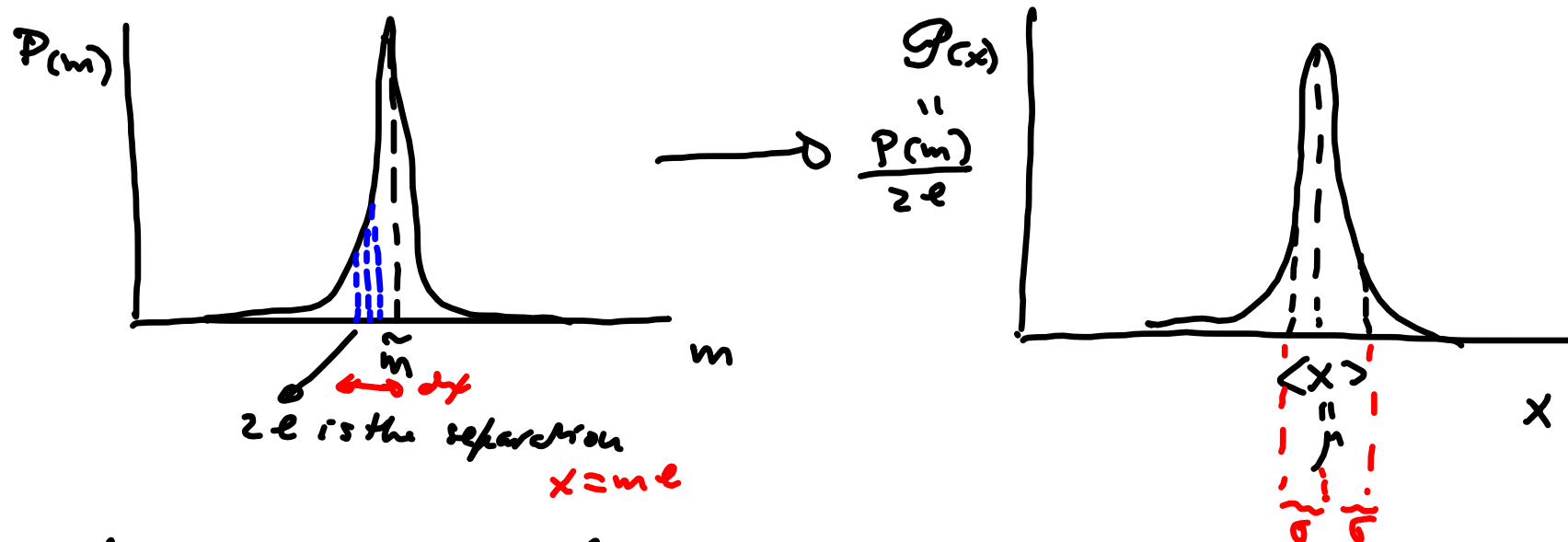
$$P(m) = W\left(\frac{N+m}{2}\right) = [2\pi Npq]^{-1/2} e^{-\frac{[m - N(p-q)]^2}{2Npq}}$$

$\langle m \rangle$   
 $2.4 Npq = 2 \langle \Delta m^2 \rangle$

$$= [2\pi Npq]^{-1/2} e^{-\frac{m - \langle m \rangle}{2 \langle \Delta m^2 \rangle}}$$

We can express  $P(m)$  in terms of the actual displacement  $x = m\ell$        $\ell$ : size of the step.

If  $\ell \ll L$        $L$ : length of interest such as the length of a container.  
then  $x$  can be considered continuous.       $\ell$ : of the order of the mean free path.



in each interval  $dx$   
 there are  $\frac{dx}{2l}$  values of  $m$ .

$P(x)$  is a probability density.

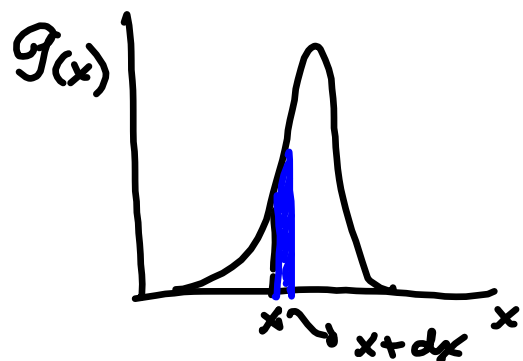
$$P(x) dx = \frac{P(m)}{2l} dx = \frac{1}{2\pi\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

Gaussian distribution

$$\mu \equiv (p-q)Nl$$

$$\sigma \equiv 2\sqrt{Npq}l$$

$P(x) dx$  is the probability of finding the walker in the interval  $(x, x+dx)$  and it is given by the blue area in the figure.



Notice that if  $p \neq q$  the gaussian approximation is still valid when  $N \rightarrow \infty$  as long as  $p$  is not too small. If  $p \ll 1$  then

$$\langle \Delta n_i^2 \rangle = \sqrt{\frac{q}{p}} \sqrt{\frac{1}{N}}$$

and it does not go to zero. You'll get **Poisson's distribution**.

Probability Distributions for more than 1  
random variables: discrete case.

$P(u_i, v_j)$

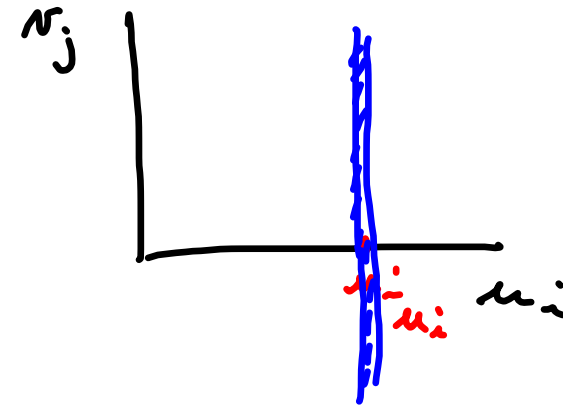
prob that a random variable  
 $u$  takes the value  $u_i$  of  
 $M$  possible values and a  
random variable  $v$  takes the  
value  $v_j$  of  $N$  possible values,

$P(u_i, v_j)$  is normalized  $i=1, \dots, M$  and  $j=1, \dots, N$ .

$$\sum_{i=1}^M \sum_{j=1}^N P(u_i, v_j) = 1$$

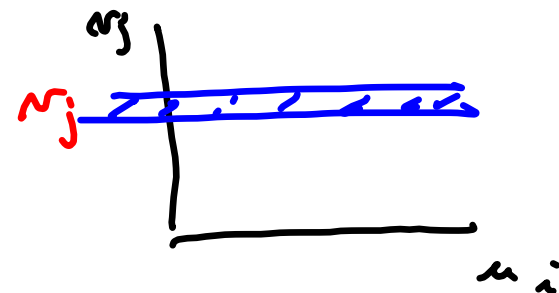
- $P_{\mu}(u_i) = \sum_{j=1}^N P(u_i, r_j) \textcircled{5}$

prob of selecting  $u_i$  regardless of the values of  $v$ .



- $P_{\sigma}(r_j) = \sum_{i=1}^M P(u_i, r_j)$

prob of selecting  $r_j$  regardless of the value of  $u$ .



- Then

$$\sum_{i=1}^N P_{\mu}(u_i) \textcircled{5} = \sum_{i=1}^N \sum_{j=1}^M P(u_i, r_j) = 1$$

- If  $u$  and  $v$  are independent variables, or uncorrelated variables then

$$P(u_i, v_j) = P_u(u_i) P_v(v_j)$$

- $\overline{F}(u, v) = \sum_{i,j} F(u_i, v_j) P(u_i, v_j)$

- $\overline{f}(u) = \sum_{i,j} f(u_i) P(u_i, v_j) = \sum_{i=1}^M f(u_i) \underbrace{\sum_{j=1}^N P(u_i, v_j)}_{P_u(u_i)}$

$$\stackrel{\textcircled{5}}{=} \sum_{i=1}^M f(u_i) P_u(u_i)$$

- $\overline{F + G} = \overline{F} + \overline{G}$

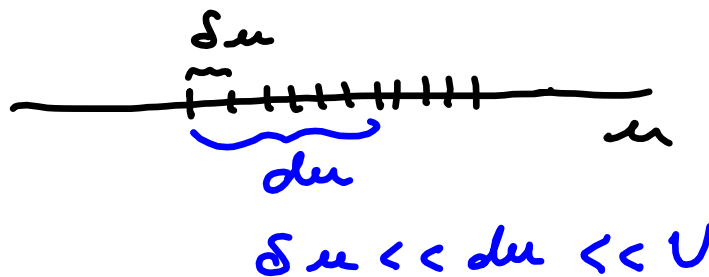
- $\overline{f(u) g(v)} = \overline{f(u)} \overline{g(v)}$

# Probability Density (continuous case)

If the length of interest is much larger than the interval between the values of the discrete random variables we can define a continuous probability density

$$\frac{P(u)}{\delta u} du = P(u) du$$

↳ probability density.

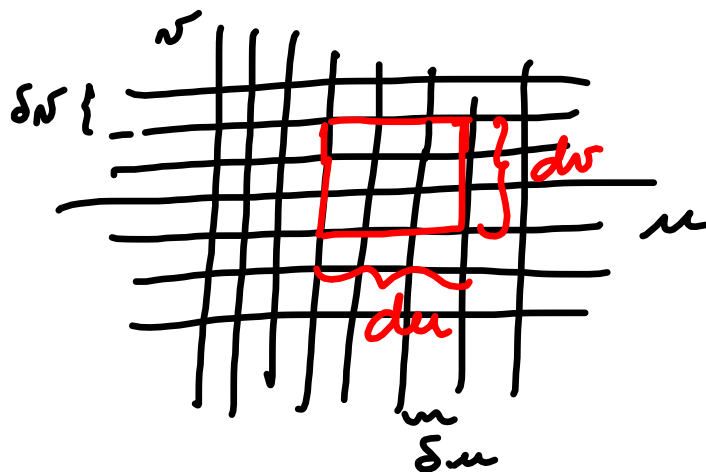


Now sums become integrals:

$$\sum_{i=1}^N P(u_i) = 1 \longrightarrow \int_{a_1}^{a_2} P(u) du = 1$$

$a_1, a_2$   
range of  $u$

Beyond 1 variable:



$$P(u, v) du dv = \frac{P(u, v) du dv}{\delta u \delta v}$$

probability that  
 $u$  and  $v$  fall  
in  $dA = du dv$  (red area)



Now

$$\int_{a_1}^{a_2} \int_{b_1}^{b_2} du dv P(u, v) = 1 \quad \text{etc.}$$